The Role of Statistics in Machine Learning

Harrison B. Prosper

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The Role of Statistics in Machine Learning

A Perspective

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Quarks To Cosmos with AI Virtual Conference: July 12-16, 2021

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We need to know when we don't know

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For the most part, the field of machine learning has developed independently of the much older field of statistics.

However, over the past decade there has been a growing recognition of the importance of quantifying the quality, reliability, or accuracy of machine learning models.

Quantifying uncertainty is especially important in physics. However, even a cursory inspection of the living review https://iml-wg.github.io/HEPML-LivingReview/ suggests that while uncertainty quantification of ML models is firmly on the stage, it has yet to reach its center.

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Most ML models are trained, that is fitted to data, $D = x_i, y_i$, $i = 1, \dots, N$, by minimizing a function that statisticians refer to as the empirical risk,

$$R(\omega) = rac{1}{N} \sum_{i=1}^{N} L(y_i, f_i), \quad f_i \equiv f(x_i, \omega),$$

where L(y, f) is a loss function and $f(x, \omega)$ is a model, which today is invariably a deep neural network with a huge number of parameters ω , such as ...

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...this one,

$f(x,\omega) = \mathsf{LSTM}(x,h,c,\omega)$

Inputs
$$(x_t, h_{t-1}, c_{t-1})$$

 $g_t = \tanh(W_{ig}x_t + b_{ig} + W_{hg}h_{t-1} + b_{hg})$
 $i_t = \sigma(W_{ii}x_t + b_{ii} + W_{hi}h_{t-1} + b_{hi})$
 $f_t = \sigma(W_{if}x_t + b_{if} + W_{hf}h_{t-1} + b_{hf})$

oututs (o_t, h_t, c_t) $o_t = \sigma(W_{io}x_t + b_{io} + W_{ho}h_{t-1} + b_{ho})$ $c_t = f_t \odot c_{t-1} + i_t \odot g_t$ $h_t = o_t \odot \tanh(c_t),$

being explored by our intrepid symbolic-AI hackathoners.

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 $R(\omega)$ is minimized using some variation of stochastic gradient descent. The goal is to find a good approximation to the minimum not of the function $R(\omega)$ but rather of the risk functional,

$$R[f] = \int \int L(y, f) p(x, y) \, dx \, dy,$$

of which $R(\omega)$ is a Monte Carlo approximation.

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$$R[f] = \int \int L(y, f) \, p(x, y) \, dx \, dy,$$

of which $R(\omega)$ is a Monte Carlo approximation.

If the model $f(x, \omega)$ has enough *capacity*, then it will be possible to set its functional derivative

$$\frac{\delta R}{\delta f} = \int \frac{\partial L}{\partial f} p(x, y) \, dy = 0,$$
$$= p(x) \int \frac{\partial L}{\partial f} p(y|x) \, dy = 0.$$

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Moreover, if

 $p(x) > 0 \quad \forall x,$

then we conclude that the equation that is being solved implicitly using stochastic gradient descent is an *approximation* to

$$\int \frac{\partial L}{\partial f} p(y|x) \, dy = 0.$$

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then we conclude that the equation that is being solved implicitly using stochastic gradient descent is an *approximation* to

$$\int \frac{\partial L}{\partial f} \, p(y|x) \, dy = 0.$$

If the approximation is good, then the fitted model will generalize well. A model generalizes well if the fitted model $f(x, \hat{\omega})$ is a good approximation to the model $f(x, \omega^*)$, which is the solution to the equation above.

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Example 1: $L(f, y) = (y - f)^2$

In 1990, it was established¹ that the quadratic loss,

$$L(y,f)=(y-f)^2,$$

when used with the discrete targets $y \in \{1, 0\}$ associated with classes C_1 and C_2 , respectively, and their associated priors $p(C_1)$ and $p(C_2)$, yields the result

$$f(x,\hat{\omega}) \approx p(C_1|x), = \frac{p(x|C_1) p(C_1)}{p(x|C_1) p(C_1) + p(x|C_2) p(C_2)}$$

which also follows from the cross entropy loss.

¹E. A. Wan, "Neural network classification: a Bayesian interpretation," in IEEE Transactions on Neural Networks, vol. 1, no. 4, pp. 303-305, 1990

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Example 2: $L(f, y) = (y - f)^2$

If the targets y come from a continuous set, then the quadratic loss yields a model that approximates

$$f(x,\hat{\omega})\approx\int y\,p(y|x)\,dy,$$

that is, the mean of the posterior density

$$p(y|x) = p(x|y) p(y)/p(x),$$

which, note, necessarily depends on the prior p(y).

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Example 3: $L(f, y) = f \delta(y - 1) + (f \log f - f) \delta(y)$

This loss function, with $p(C_1) = p(C_2)$, yields the result

$$f = \exp(-r)$$
, where
 $r = \frac{p(C_1|x)}{p(C_2|x)} = \frac{p(x|C_1)}{p(x|C_2)}$,

that is, the negative log of the ML model approximates the likelihood ratio.

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that is, the negative log of the ML model approximates the likelihood ratio.

The point here is that every ML model that uses the *same* training data and is fitted with the *same* loss function will approximate the *same* mathematical function.

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Since 2010, more than 2,500 papers have been published on uncertainty quantification, though few by particle physicists.

The most recent papers seem to be converging on the use of the Bayesian approach to quantify both aleatoric (i.e., statistical) as well as epistemic (i.e., model) uncertainty.

The Bayesian approach is conceptually simple: given a model $f(x, \omega)$, one computes the posterior density $p(\omega|D)$ of its parameters.

From $p(\omega|D)$ different measures of uncertainty can be computed for any quantity that depends on ω .

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The Bayesian approach is conceptually simple: given a model $f(x, \omega)$, one computes the posterior density $p(\omega|D)$ of its parameters.

From $p(\omega|D)$ different measures of uncertainty can be computed for any quantity that depends on ω .

The devil, as always, is in the details and the computational burden...

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The posterior density of the model parameters is given by Bayes' theorem,

$$p(\omega|D) = rac{p(D,\omega)}{p(D)} = rac{p(Y|X,\omega) p(X|\omega) \pi(\omega)}{p(Y|X) p(X)},$$

where $D \equiv X, Y$ and X and Y represent all values of x and y, respectively. The function $\pi(\omega)$ is a prior density over the model parameter space.

In practice, the training data X are independent of the model parameters; therefore, $p(X|\omega) = p(X)$ and we can write

$$p(\omega|D) = rac{p(Y|X,\omega) \pi(\omega)}{p(Y|X)}$$

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For relatively small models it is possible to represent the posterior density, $p(\omega|D)$, as a point cloud sampled using, for example, a Markov chain Monte Carlo (MCMC) method such as Hamiltonian Monte Carlo (HMC) or a variant.

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Unfortunately, for large models MCMC is too slow and the trend now is to focus on tractable approximations $q(\omega|\theta)$ of $p(\omega|D)$ from which it is easy, and fast, to sample points ω_k .

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Unfortunately, for large models MCMC is too slow and the trend now is to focus on tractable approximations $q(\omega|\theta)$ of $p(\omega|D)$ from which it is easy, and fast, to sample points ω_k .

The parameters θ are chosen to get the best match between $q(\omega|\theta)$ and $p(\omega|D)$, in principle by minimizing the Kullback-Leibler (KL) divergence

$$\mathsf{KL}(q|p) = \int q(\omega| heta) \log\left[rac{q(\omega| heta)}{p(\omega|D)}
ight] \, d\omega,$$

between them.

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Alas we can't minimize

$$\mathsf{KL}(q|p) = \int q(\omega| heta) \log\left[rac{q(\omega| heta)}{p(\omega|D)}
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because it depends on the very thing, $p(\omega|D)$, we wish to approximate!

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Alas we can't minimize

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because it depends on the very thing, $p(\omega|D)$, we wish to approximate!

But, we can write the KL divergence as

$$egin{aligned} \mathsf{KL}(q|p) &= \log p(Y|X) \ &+ \mathsf{KL}(q|\pi) - \int q(\omega| heta) \, \log p(Y|X,\omega) \, d\omega. \end{aligned}$$

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The negative of the term in blue is called the the evidence lower bound (ELBO), which contains *known* quantities.

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Therefore, optimizing the function

$$\mathsf{ELBO}(heta) = \mathsf{KL}(q|\pi) - \int q(\omega|\theta) \log p(Y|X,\omega) \, d\omega,$$

with respect to θ to find the best fit of $q(\omega|\theta)$ to the posterior density $p(\omega|D)$ is feasible, in principle, and, crucially, it is equivalent to optimizing the original KL divergence. But, alas, this is very challenging for large models and data sets.

²Yarin Gal and Zoubin Ghahramani, arXiv:1506.02142v6, 2016 = $9 \circ \circ$

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In 2016, Gal and Ghahramani² introduced an astonishingly simple method to quantify the uncertainty of any ML model, $f(x, z, \omega)$, trained using dropout random variables z before every layer.

²Yarin Gal and Zoubin Ghahramani, arXiv:1506.02142v6, 2016 📃 🔗 < 🔿

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Gal and Ghahramani showed that if the dropout variables are used *after* training then M forward passes through a model for given input data x,

 $\hat{y}_m = f(x, z_m, \hat{\omega}), \quad m = 1, \cdots, M,$

yields a sample $\{\hat{y}_m\}$ that is a point cloud approximation to the posterior density $p(\omega|D)$.

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 $\hat{y}_m = f(x, z_m, \hat{\omega}), \quad m = 1, \cdots, M,$

yields a sample $\{\hat{y}_m\}$ that is a point cloud approximation to the posterior density $p(\omega|D)$.

Specifically, the authors show that MC dropout is mathematically equivalent to approximating $p(\omega|D)$ using a function $q(\omega|\theta)$ constructed from a Gaussian process (GP).

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Monte Carlo dropout as a Gaussian process

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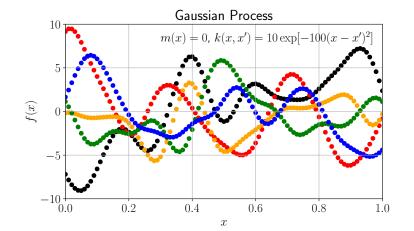
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A GP is a multivariate Gaussian with an infinite number of random variables and is characterized by two functions

m(x): mean function, k(x, x') : covariance function.



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A *GP* can be used to approximate a probability density p(f|D) = p(D|f) GP(f) over the space of functions.

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The problem with all current methods is that the uncertainty estimates are not calibrated.

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The problem with all current methods is that the uncertainty estimates are not calibrated.

If a 68% credible interval is computed from a Bayesian method how reliable is it?

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The problem with all current methods is that the uncertainty estimates are not calibrated.

- If a 68% credible interval is computed from a Bayesian method how reliable is it?
- With enough deep thinking one could, in principle, arrive at a prior $\pi(\omega)$ that would yield credible intervals that are reliable from a Bayesian viewpoint.

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The problem with all current methods is that the uncertainty estimates are not calibrated.

- If a 68% credible interval is computed from a Bayesian method how reliable is it?
- With enough deep thinking one could, in principle, arrive at a prior π(ω) that would yield credible intervals that are reliable from a Bayesian viewpoint.
- However, most physicists would not find the credible intervals acceptable if they deviated greatly from 68% confidence intervals.

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- If a 68% credible interval is computed from a Bayesian method how reliable is it?
- With enough deep thinking one could, in principle, arrive at a prior π(ω) that would yield credible intervals that are reliable from a Bayesian viewpoint.
- However, most physicists would not find the credible intervals acceptable if they deviated greatly from 68% confidence intervals.

WANTED: *Feasible* and *general* ways to check the coverage of ML-based intervals, credible or otherwise.

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It widely recognized that uncertainty quantification of ML models is crucial and will become more so as we become more dependent on them.

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- It widely recognized that uncertainty quantification of ML models is crucial and will become more so as we become more dependent on them.
- The machine learning/AI field seems to be converging on the use of Bayesian methods to quantify uncertainty.

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- However, these methods need to be calibrated.

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- It widely recognized that uncertainty quantification of ML models is crucial and will become more so as we become more dependent on them.
- The machine learning/AI field seems to be converging on the use of Bayesian methods to quantify uncertainty.
- However, these methods need to be calibrated.
- For many physicists, and many statisticians, the preferred way to do so is to calibrate *all* proposed intervals using frequentist methods.

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- The machine learning/AI field seems to be converging on the use of Bayesian methods to quantify uncertainty.
- However, these methods need to be calibrated.
- For many physicists, and many statisticians, the preferred way to do so is to calibrate *all* proposed intervals using frequentist methods.
- The problem is how to do this in practice.

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